

Likelihood calculations for *vs_n*

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1 Introduction

This vignette contains the computations that underlie the numerical code of *vs_n*. If you are a new user and looking for an introduction on how to **use** *vs_n*, please refer to the vignette *Robust calibration and variance stabilization with vs_n*, which is provided separately.

2 Setup and Notation

Consider the model

$$\operatorname{arsinh}(f(b_i) \cdot y_{ki} + a_i) = \mu_k + \varepsilon_{ki} \quad (1)$$

where μ_k , for $k = 1, \dots, n$, and a_i , b_i , for $i = 1, \dots, d$ are real-valued parameters, f is a function $\mathbb{R} \rightarrow \mathbb{R}$ (see below), and ε_{ki} are i.i.d. Normal with mean 0 and variance σ^2 . y_{ki} are the data. In applications to `μarray` data, k indexes the features and i the arrays and/or colour channels.

Examples for f are $f(b) = b$ and $f(b) = e^b$. The former is the most obvious choice; in that case we will usually need to require $b_i > 0$. The choice $f(b) = e^b$ assures that the factor in front of y_{ki} is positive for all $b \in \mathbb{R}$, and as it turns out, simplifies some of the computations.

In the following calculations, I will also use the notation

$$Y \equiv Y(y, a, b) = f(b) \cdot y + a \quad (2)$$

$$h \equiv h(y, a, b) = \operatorname{arsinh}(f(b) \cdot y + a). \quad (3)$$

The probability of the data $(y_{ki})_{k=1\dots n, i=1\dots d}$ lying in a certain volume element of y -space (hyperrectangle with sides $[y_{ki}^\alpha, y_{ki}^\beta]$) is

$$P = \prod_{k=1}^n \prod_{i=1}^d \int_{y_{ki}^\alpha}^{y_{ki}^\beta} dy_{ki} \, p_{\text{Normal}}(h(y_{ki}), \mu_k, \sigma^2) \frac{dh}{dy}(y_{ki}), \quad (4)$$

where μ_k is the expectation value for feature k and σ^2 the variance.

With

$$p_{\text{Normal}}(x, \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right) \quad (5)$$

the likelihood is

$$L = \left(\frac{1}{\sqrt{2\pi\sigma^2}}\right)^{nd} \prod_{k=1}^n \prod_{i=1}^d \exp\left(-\frac{(h(y_{ki}) - \mu_k)^2}{2\sigma^2}\right) \cdot \frac{dh}{dy}(y_{ki}). \quad (6)$$

For the following, I will need the derivatives

$$\frac{\partial Y}{\partial a} = 1 \quad (7)$$

$$\frac{\partial Y}{\partial b} = y \cdot f'(b) \quad (8)$$

$$\frac{dh}{dy} = \frac{f(b)}{\sqrt{1 + (f(b)y + a)^2}} = \frac{f(b)}{\sqrt{1 + Y^2}}, \quad (9)$$

$$\frac{\partial h}{\partial a} = \frac{1}{\sqrt{1 + Y^2}}, \quad (10)$$

$$\frac{\partial h}{\partial b} = \frac{y}{\sqrt{1 + Y^2}} \cdot f'(b). \quad (11)$$

Note that for $f(b) = b$, we have $f'(b) = 1$, and for $f(b) = e^b$, $f'(b) = f(b) = e^b$.

3 Likelihood for Incremental Normalization

Here, *incremental normalization* means that the model parameters μ_1, \dots, μ_n and σ^2 are already known from a fit to a previous set of μ arrays, i.e. a set of reference arrays. See Section 4 for the profile likelihood approach that is used if μ_1, \dots, μ_n and σ^2 are not

known and need to be estimated from the same data. Versions ≥ 2.0 of the *vsu* package implement both of these approaches; in versions 1.X only the profile likelihood approach was implemented, and it was described in the initial publication [1].

First, let us note that the likelihood (6) is simply a product of independent terms for different i . We can optimize the parameters (a_i, b_i) separately for each $i = 1, \dots, d$. From the likelihood (6) we get the i -th negative log-likelihood

$$-\log(L) = \sum_{i=1}^d -LL_i \quad (12)$$

$$-LL_i = \frac{n}{2} \log(2\pi\sigma^2) + \sum_{k=1}^n \left(\frac{(h(y_{ki}) - \mu_k)^2}{2\sigma^2} + \log \frac{\sqrt{1 + Y_{ki}^2}}{f(b_i)} \right) \quad (13)$$

$$= \frac{n}{2} \log(2\pi\sigma^2) - n \log f(b_i) + \sum_{k=1}^n \left(\frac{(h(y_{ki}) - \mu_k)^2}{2\sigma^2} + \frac{1}{2} \log(1 + Y_{ki}^2) \right) \quad (14)$$

This is what we want to optimize as a function of a_i and b_i . The optimizer benefits from the derivatives. The derivative with respect to a_i is

$$\begin{aligned} \frac{\partial}{\partial a_i}(-LL_i) &= \sum_{k=1}^n \left(\frac{h(y_{ki}) - \mu_k}{\sigma^2} + \frac{Y_{ki}}{\sqrt{1 + Y_{ki}^2}} \right) \cdot \frac{1}{\sqrt{1 + Y_{ki}^2}} \\ &= \sum_{k=1}^n \left(\frac{r_{ki}}{\sigma^2} + A_{ki} Y_{ki} \right) A_{ki} \end{aligned} \quad (15)$$

and with respect to b_i

$$\begin{aligned} \frac{\partial}{\partial b_i}(-LL_i) &= -n \frac{f'(b_i)}{f(b_i)} + \sum_{k=1}^n \left(\frac{h(y_{ki}) - \mu_k}{\sigma^2} + \frac{Y_{ki}}{\sqrt{1 + Y_{ki}^2}} \right) \cdot \frac{y_{ki}}{\sqrt{1 + Y_{ki}^2}} \cdot f'(b_i) \\ &= -n \frac{f'(b_i)}{f(b_i)} + f'(b_i) \sum_{k=1}^n \left(\frac{r_{ki}}{\sigma^2} + A_{ki} Y_{ki} \right) A_{ki} y_{ki} \end{aligned} \quad (16)$$

Here, I have introduced the following shorthand notation for the “intermediate results” terms

$$r_{ki} = h(y_{ki}) - \mu_k \quad (17)$$

$$A_{ki} = \frac{1}{\sqrt{1 + Y_{ki}^2}}. \quad (18)$$

Variables for these intermediate values are also used in the C code to organise the computations of the gradient.

4 Profile Likelihood

If μ_1, \dots, μ_n and σ^2 are not already known, we can plug in their maximum likelihood estimates, obtained from optimizing LL for μ_1, \dots, μ_n and σ^2 :

$$\hat{\mu}_k = \frac{1}{d} \sum_{j=1}^d h(y_{kj}) \quad (19)$$

$$\hat{\sigma}^2 = \frac{1}{nd} \sum_{k=1}^n \sum_{j=1}^d (h(y_{kj}) - \hat{\mu}_k)^2 \quad (20)$$

into the negative log-likelihood. The result is called the negative profile log-likelihood

$$-PLL = \frac{nd}{2} \log(2\pi\hat{\sigma}^2) + \frac{nd}{2} - n \sum_{j=1}^d \log f(b_j) + \frac{1}{2} \sum_{k=1}^n \sum_{j=1}^d \log \sqrt{1 + Y_{kj}^2}. \quad (21)$$

Note that this no longer decomposes into a sum of terms for each j that are independent of each other – the terms for different j are coupled through Equations (19) and (20). We need the following derivatives.

$$\begin{aligned} \frac{\partial \hat{\sigma}^2}{\partial a_i} &= \frac{2}{nd} \sum_{k=1}^n r_{ki} \frac{\partial h(y_{ki})}{\partial a_i} \\ &= \frac{2}{nd} \sum_{k=1}^n r_{ki} A_{ki} \end{aligned} \quad (22)$$

$$\frac{\partial \hat{\sigma}^2}{\partial b_i} = \frac{2}{nd} \cdot f'(b_i) \sum_{k=1}^n r_{ki} A_{ki} y_{ki} \quad (23)$$

So, finally

$$\begin{aligned} \frac{\partial}{\partial a_i}(-PLL) &= \frac{nd}{2\hat{\sigma}^2} \cdot \frac{\partial \hat{\sigma}^2}{\partial a_i} + \sum_{k=1}^n A_{ki}^2 Y_{ki} \\ &= \sum_{k=1}^n \left(\frac{r_{ki}}{\hat{\sigma}^2} + A_{ki} Y_{ki} \right) A_{ki} \end{aligned} \quad (24)$$

$$\frac{\partial}{\partial b_i}(-PLL) = -n \frac{f'(b_i)}{f(b_i)} + f'(b_i) \sum_{k=1}^n \left(\frac{r_{ki}}{\hat{\sigma}^2} + A_{ki} Y_{ki} \right) A_{ki} y_{ki} \quad (25)$$

5 Summary

Likelihoods, from Equations (12) and (21):

$$-LL_i = \underbrace{\frac{n}{2} \log(2\pi\sigma^2)}_{\text{scale}} + \underbrace{\sum_{k=1}^n \frac{(h(y_{ki}) - \mu_k)^2}{2\sigma^2}}_{\text{residuals}} \underbrace{-n \log f(b_i) + \frac{1}{2} \sum_{k=1}^n \log(1 + Y_{ki}^2)}_{\text{jacobian}} \quad (26)$$

$$-PLL = \underbrace{\frac{nd}{2} \log(2\pi\hat{\sigma}^2)}_{\text{scale}} + \underbrace{\frac{nd}{2}}_{\text{residuals}} + \underbrace{\sum_{i=1}^d \left(-n \log f(b_i) + \frac{1}{2} \sum_{k=1}^n \log(1 + Y_{ki}^2) \right)}_{\text{jacobian}} \quad (27)$$

The computations in the C code are organised into steps for computing the terms “scale”, “residuals” and “jacobian”.

Partial derivatives with respect to a_i , from Equations (15) and (24):

$$\frac{\partial}{\partial a_i}(-LL_i) = \sum_{k=1}^n \left(\frac{r_{ki}}{\sigma^2} + A_{ki} Y_{ki} \right) A_{ki} \quad (28)$$

$$\frac{\partial}{\partial a_i}(-PLL) = \sum_{k=1}^n \left(\frac{r_{ki}}{\hat{\sigma}^2} + A_{ki} Y_{ki} \right) A_{ki} \quad (29)$$

Partial derivatives with respect to b_i , from Equations (16) and (25):

$$\frac{\partial}{\partial b_i}(-LL_i) = -n \frac{f'(b_i)}{f(b_i)} + f'(b_i) \sum_{k=1}^n \left(\frac{r_{ki}}{\sigma^2} + A_{ki} Y_{ki} \right) A_{ki} y_{ki} \quad (30)$$

$$\frac{\partial}{\partial b_i}(-PLL) = -n \frac{f'(b_i)}{f(b_i)} + f'(b_i) \sum_{k=1}^n \left(\frac{r_{ki}}{\hat{\sigma}^2} + A_{ki} Y_{ki} \right) A_{ki} y_{ki}. \quad (31)$$

Note that the terms have many similarities – this is used in the implementation in the C code.

References

- [1] W. Huber, A. von Heydebreck, H. Sltmann, A. Poustka, and M. Vingron. Variance stablization applied to microarray data calibration and to quantification of differential expression. *Bioinformatics*, 18:S96–S104, 2002. 2
- [2] W. Huber, A. von Heydebreck, H. Sltmann, A. Poustka, and M. Vingron. Parameter estimation for the calibration and variance stabilization of microarray data. *Statistical Applications in Genetics and Molecular Biology*, Vol. 2: No. 1, Article 3, 2003. <http://www.bepress.com/sagmb/vol2/iss1/art3>